# **Coordinate-Update Algorithms can Efficiently Detect Infeasible Optimization Problems**

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Introduction

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### **Definition (Fixed Point Iteration)**

For a given operator  ${\mathbb T},$  we call an iterative method

$$x^{k+1} = \mathbb{T}x^k, \quad k = 0, 1, 2, \cdots$$

as Fixed Point Iteration, or FPI.

#### Theorem (Banach-Fixed Point Theorem)

When (X, d) is a non-empty complete metric space and an operator  $\mathbb{T} : X \to X$  is L-Lipschitz with L < 1, then

- there exists an unique fixed point of  $\mathbb{T}$ ,  $x^* = \mathbb{T}x^*$ ,
- FPI with  $\mathbb{T}$  converges to  $x^*$ ,  $\lim_{k\to\infty} x^k = x^*$ .

# **Fixed Point Iteration**

**Remark.** However, in most of *FPI*-based convex optimization solvers, the operator is only guaranteed to be a 1-Lipschitz.

### **Definition** ( $\theta$ -averaged operator)

An operator  ${\mathbb T}$  is  $\theta\text{-averaged}$  if it can be described as

 $\mathbb{T} = (1 - \theta)\mathbb{I} + \theta\mathbb{C}, \quad \theta \in (0, 1], \quad \mathbb{C} \text{ is 1-Lipschitz.}$ 

The set of fixed points  $\operatorname{Fix} \mathbb{T}$  coincides with  $\operatorname{Fix} \mathbb{C}$ .

#### Theorem (Averaged Fixed Point Theorem)

When  $\mathcal{H}$  is a non-empty Real Hilbert space and an operator  $\mathbb{T} : \mathcal{H} \to \mathcal{H}$  is  $\theta$ -averaged with  $\theta \in (0, 1)$  and has a nonempty fixed point,  $\operatorname{Fix} \mathbb{T} \neq \emptyset$ , then

$$\lim_{k\to\infty} x^k = x^*, \quad x^* \in \operatorname{Fix} \mathbb{T}.$$

# Fixed Point Iteration - Inconsistent case

**Remark.** Fix  $\mathbb{T} = \emptyset$  is possible when  $\mathbb{T}$  is 1-Lipschitz or  $\theta$ -averaged mapping. For example, consider a translation mapping.

#### Theorem (Pazy, 1971)

When  $\mathcal{H}$  is a non-empty Real Hilbert space and an operator  $\mathbb{T}: \mathcal{H} \to \mathcal{H}$  is 1-Lipschitz mapping, then

$$\lim_{k\to\infty}\frac{x^k}{k}=-\mathbf{v},$$

where  $\mathbf{v}$  is a minimal norm vector of closed convex set range  $(\mathbb{I} - \mathbb{T})$ .

- We call  $\mathbf{v}$  as an *infimal displacement vector* of  $\mathbb{T}$ .
- We call  $x^k/k$  as a *normalized iterate* of *FPI* by  $\mathbb{T}$ .

**Remark.** If Fix  $\mathbb{T} \neq \emptyset$ , then  $\mathbf{v} = 0$ .

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### Definition (Classical RC-FPI)

Partition  $\mathbb{R}^n$  into a *m* block coordinates and write a vector in  $\mathbb{R}^n$  as

$$x=(x_1,x_2,\cdots,x_m).$$

Consider an operator  $\mathbb{T}$  on  $\mathbb{R}^n$  and define a *randomized operator*  $\mathbb{T}_i$  as:

 $\mathbb{T}x = ((\mathbb{T}x)_1, (\mathbb{T}x)_2, \cdots, (\mathbb{T}x)_m), \quad \mathbb{T}_i x = (x_1, x_2, \cdots, (\mathbb{T}x)_i, \cdots, x_m).$ 

A randomized update by  $\mathbb{T}_{i^k}$  with IID  $i^k$ 's is called *RC-FPI*, short for Randomized Coordinate Fixed Point Iteration:

$$x^{k+1} = \mathbb{T}_{i^k} x^k, \quad i^0, i^1, \cdots \stackrel{iid}{\sim} \mathsf{Uniform}(m).$$

### Theorem (Convergence of RC-FPI)

When an operator  $\mathbb{T} : \mathbb{R}^n \to \mathbb{R}^n$  is  $\theta$ -averaged with  $\theta \in (0, 1)$  and has a nonempty fixed point,  $\operatorname{Fix} \mathbb{T} \neq \emptyset$ , then with probability 1,

$$\lim_{k\to\infty} x^k = x^*, \quad x^* \in \operatorname{Fix} \mathbb{T}.$$

**Remark.** Many optimization solvers use (RC-FPI) due to its faster convergence speed. Faster speed is not guaranteed but it is not slower if the operator is *coordinate friendly*.

Question. What is the asymptotic behavior of RC-FPI when  $Fix T = \emptyset$ ? Answers. Here's the results this work has found for the first time.

• Convergence result analogous of Pazy's work, both in  $L^2$  and a.s.:

$$\lim_{k\to\infty}\frac{x^k}{k}=-\alpha\mathbf{v}.$$

• Upper bound of the variance analogous to the CLT:

$$\limsup_{k \to \infty} k \operatorname{Var}\left(\frac{x^k}{k}\right) \le \left(\alpha - \alpha^2\right) \|\mathbf{v}\|^2.$$

Here,  $\alpha$  is a probability of an update for each coordinate.

### Main Goal of the Paper



Example of (RC-FPI), 100 iterations, ran 100 times.

Let's extend RC-FPI to more general setting.

### Underlying space:

• The underlying space is a real Hilbert space  $\mathcal{H}$ , consisted of *m* real Hilbert spaces.

$$\mathcal{H}=\mathcal{H}_1\oplus\mathcal{H}_2\oplus\ldots\mathcal{H}_m.$$

• An element  $u \in \mathcal{H}$  can be decomposed into m blocks as

$$u = (u_1, u_2, \ldots, u_m), \quad u_i \in \mathcal{H}_i,$$

and  $u_i$  is called the *i*th block coordinates of u.

# **RC-FPI setting: Underlying space**

 $\bullet$  The Hilbert space  ${\cal H}$  has its induced norm and inner product as

$$\|x\|^2 = \sum_{i=1}^m \|x_i\|_i^2, \qquad \langle x, y \rangle = \sum_{i=1}^m \langle x_i, y_i \rangle_i,$$

for all  $x, y \in \mathcal{H}$ , where  $\|\cdot\|_i$  and  $\langle \cdot, \cdot \rangle_i$  are from  $\mathcal{H}_i$ .

Consider a linear, bounded, self-adjoint and positive definite operator
 M : H → H. The M-norm and M-inner product of H are defined as

$$\|x\|_M = \sqrt{\langle x, Mx \rangle}, \qquad \langle x, y \rangle_M = \langle x, My \rangle,$$

• The *M*-variance of a random variable *X* with the domain  $\mathcal{H}$  as

$$\operatorname{Var}_{M}[X] = \mathbb{E}[||X||_{M}^{2}] - ||\mathbb{E}[X]||_{M}^{2}.$$

**Remark.** When an operator  $\mathbb{C} : \mathcal{H} \to \mathcal{H}$  is 1-Lipschitz of a real Hilbert space  $\mathcal{H}$ , then  $S = \mathbb{I} - \mathbb{C}$  is (1/2)-coccersive operator:

$$\langle x-y, \mathbf{S}x-\mathbf{S}y \rangle \geq \frac{1}{2} \|\mathbf{S}x-\mathbf{S}y\|^2, \quad \forall x, y \in \mathcal{H}.$$

Consider a  $\theta$ -averaged  $\mathbb{T} = (1 - \theta)\mathbb{I} + \theta\mathbb{C}$ , then we can rewrite  $\mathbb{T}$  as

 $\mathbb{T} = \mathbb{I} - \theta \mathbb{S}$ ,  $\mathbb{S}$  is (1/2)-cocoersive.

For (1/2)-cocoercive operator  $S = \theta^{-1}(\mathbb{I} - \mathbb{T})$ , define  $S_i : \mathcal{H} \to \mathcal{H}$  as:

$$\mathbf{S} = \sum_{i=1}^{m} \mathbf{S}_i, \quad \mathbf{S}_i : \mathcal{H} \to \mathbf{0} \times \mathbf{0} \times \cdots \times \mathcal{H}_i \times \cdots \times \mathbf{0}.$$

**Randomized operator:** Consider a  $\theta$ -averaged operator  $\mathbb{T} \colon \mathcal{H} \to \mathcal{H}$ .

• For a selection vector  $\mathcal{I} \in [0,1]^m$ , define  $\mathbb{S}_{\mathcal{I}}, \mathbb{T}_{\mathcal{I}} \colon \mathcal{H} \to \mathcal{H}$  as:

$$\mathbf{S}_{\mathcal{I}} = \sum_{i=1}^{m} \mathcal{I}_i \mathbf{S}_i, \qquad \mathbf{T}_{\mathcal{I}} = \mathbf{I} - \theta \mathbf{S}_{\mathcal{I}}.$$

**Randomized vector:** Similarly, for  $u \in \mathcal{H}$ , define  $u_{\mathcal{I}}$  as:

$$u_{\mathcal{I}} = \sum_{i=1}^m \mathcal{I}_i u_i, \quad u = \sum_{i=1}^m u_i, \quad u_i \in 0 \times 0 \times \cdots \times \mathcal{H}_i \times \cdots \times 0.$$

### **Coefficients:**

•  $\alpha$ : expected amount of update in each coordinate

Definition (Uniform expected step-size condition)

Uniform expected step-size condition is satisfied when  $\mathcal{I}$  is randomly sampled from a distribution on  $[0, 1]^m$  that satisfies:

$$\mathbb{E}_{\mathcal{I}}\left[\mathcal{I}\right] = \alpha \mathbf{1}.$$

•  $\beta:$  depends on the distribution of  ${\mathcal I}$  and the choice of the norm.

$$\mathbb{E}_{\mathcal{I}}\left[\left\|u_{\mathcal{I}}\right\|_{M}^{2}\right] \leq \beta \left\|u\right\|_{M}^{2}, \qquad \forall u \in \mathcal{H}.$$

#### Lemma

Let's consider the norm of  $\mathcal{H}$  as  $\|\cdot\|$ -norm, i.e. M is the identity map. If  $\mathcal{I}$  satisfies the uniform expected step-size condition, then we can say:

 $\beta \leq \alpha.$ 

Thus, we can choose  $\alpha$  as a value of  $\beta$ .

**Remark.**  $\beta$  always satisfies  $\beta \ge \alpha^2$  regardless of the choice of *M*-norm. However, smaller  $\beta$  is preferred.

**Remark.** We develop the theory with the general *M*-norm so that the theory can be extended to non-orthogonal basis.

# **RC-FPI** setting

### Definition (RC-FPI)

The randomized coordinate fixed-point iteration (RC-FPI) is:

$$x^{k+1} = \mathbb{T}_{\mathcal{I}^k} x^k, \qquad k = 0, 1, 2, \dots,$$
(RC-FPI)

where  $\mathcal{I}^0, \mathcal{I}^1, \ldots$  is sampled IID and  $x^0 \in \mathcal{H}$  is a starting point.

**Remark.** With uniform expected step-size condition, define  $\overline{\mathbb{T}} \colon \mathcal{H} \to \mathcal{H}$ :

$$\bar{\mathbb{T}}x = \mathbb{E}_{\mathcal{I}}\left[\mathbb{T}_{\mathcal{I}}x\right], \qquad \forall x \in \mathcal{H}.$$

Equivalently,  $\overline{\mathbb{T}} = \mathbb{I} - \alpha \theta \mathbb{S}$ . We can define a FPI by  $\overline{\mathbb{T}}$ :

$$z^{k+1} = \overline{\mathbb{T}} z^k, \qquad k = 0, 1, 2, \dots$$
 (FPI by  $\overline{\mathbb{T}}$ )

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**Goal.** The normalized iterate of (RC-FPI) converges to  $-\alpha \mathbf{v}$ :

$$\frac{x^k}{k} \stackrel{L^2}{\to} -\alpha \mathbf{v}, \quad \frac{x^k}{k} \stackrel{\text{a.s.}}{\to} -\alpha \mathbf{v}.$$

both in  $L^2$  and almost surely.

Here's an inequality we will use repeatedly throughout the proofs.

#### Lemma

#### Consider the situation:

- $\mathbb{T}: \mathcal{H} \to \mathcal{H}$  be  $\theta$ -averaged respect to  $\|\cdot\|_M$  with  $\theta \in (0, 1]$ .
- ${\mathcal I}$  satisfies the uniform expected step-size condition.

Then for any  $x, z \in \mathcal{H}$ ,

$$\begin{split} \mathbb{E}_{\mathcal{I}} \left[ \left\| \mathbb{T}_{\mathcal{I}} x - \bar{\mathbb{T}} z \right\|_{M}^{2} \right] &\leq \left\| x - z \right\|_{M}^{2} \\ &+ \theta^{2} \left( \beta - \alpha^{2} \right) \left\| \mathbf{S} x \right\|_{M}^{2} - \alpha \theta \left( 1 - \alpha \theta \right) \left\| \mathbf{S} x - \mathbf{S} z \right\|_{M}^{2} . \end{split}$$

Notation. We will refer this inequality as One-step inequality.

# Inequality on expectation of randomized operator

### Proof.

First, substitute  $\mathbb{T}_{\mathcal{I}} = \mathbb{I} - \theta \mathbb{S}_{\mathcal{I}}$  and  $\overline{\mathbb{T}} = \mathbb{I} - \alpha \theta \mathbb{S}$  at  $\mathbb{E}_{\mathcal{I}} \left[ \left\| \mathbb{T}_{\mathcal{I}} x - \overline{\mathbb{T}} z \right\|_{M}^{2} \right]$ . Then, use (1/2)-cocoercive property of the operator  $\mathbb{S}$ .

$$\mathbb{E}\left[\left\|\mathbf{T}_{\mathcal{I}}x - \bar{\mathbf{T}}z\right\|_{M}^{2}\right]$$
  
=  $\|x - z\|_{M}^{2} + \theta^{2}\mathbb{E}\left[\left\|\mathbf{S}_{\mathcal{I}}x - \alpha\mathbf{S}z\right\|_{M}^{2}\right] - 2\alpha\theta \langle x - z, \mathbf{S}x - \mathbf{S}z \rangle_{M}$   
 $\leq \|x - z\|_{M}^{2} + \theta^{2}\mathbb{E}\left[\left\|\mathbf{S}_{\mathcal{I}}x - \alpha\mathbf{S}z\right\|_{M}^{2}\right] - \alpha\theta \|\mathbf{S}x - \mathbf{S}z\|_{M}^{2}.$ 

Finally, apply the following inequality to conclude the proof.

$$\mathbb{E}\left[\left\|\mathbf{S}_{\mathcal{I}}\mathbf{x} - \alpha\mathbf{S}\mathbf{z}\right\|_{M}^{2}\right] = \mathbb{E}\left[\left\|\left(\mathbf{S}_{\mathcal{I}}\mathbf{x} - \alpha\mathbf{S}\mathbf{x}\right) + \alpha\left(\mathbf{S}\mathbf{x} - \mathbf{S}\mathbf{z}\right)\right\|_{M}^{2}\right]\right]$$
$$\leq \left(\beta - \alpha^{2}\right)\left\|\mathbf{S}\mathbf{x}\right\|_{M}^{2} + \alpha^{2}\left\|\mathbf{S}\mathbf{x} - \mathbf{S}\mathbf{z}\right\|_{M}^{2}.$$

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### Theorem ( $L^2$ convergence of normalized iterate)

Let  $\mathbb{T}: \mathcal{H} \to \mathcal{H}$  be  $\theta$ -averaged with respect to  $\|\cdot\|$ -norm with  $\theta \in (0, 1]$ . Assume  $\mathcal{I}^0, \mathcal{I}^1, \ldots$  is sampled IID from a distribution satisfying the uniform expected step-size condition. Let  $x^0, x^1, x^2, \ldots$  be the iterates of (RC-FPI). Then

$$\frac{x^k}{k} \stackrel{L^2}{\to} -\alpha x$$

as  $k \to \infty$ , where **v** is the infimal displacement vector of  $\mathbb{T}$ .

**Remark.** We will prove for the *M*-norm with the assumption:

$$\alpha \geq \theta\beta.$$

# **Proof of** *L*<sup>2</sup> **convergence**

As mentioned, we will alternatively prove the following result.

#### Lemma

Let  $\mathbb{T}: \mathcal{H} \to \mathcal{H}$  be  $\theta$ -averaged with respect to  $\|\cdot\|_M$  with  $\theta \in (0, 1]$ . Assume  $\mathcal{I}^0, \mathcal{I}^1, \ldots$  is sampled IID from a distribution satisfying the uniform expected step-size condition and  $\beta$  satisfies that  $\beta \leq \alpha/\theta$ . Let  $x^0, x^1, x^2, \ldots$  be the iterates of (RC-FPI) and let  $z^0, z^1, z^2, \ldots$  be the iterates of (FPI by  $\overline{\mathbb{T}}$ ). Then,

$$\mathbb{E}\left[\left\|\frac{x^{k}}{k}-\frac{z^{k}}{k}\right\|_{M}^{2}\right] \leq \frac{1}{k^{2}}\left\|x^{0}-z^{0}\right\|_{M}^{2}+\frac{1}{k}A,$$

where  $\boldsymbol{v}$  is the infimal displacement vector of  $\mathbb T$  and A is a constant

$$A = (1 - \alpha \theta) \left[ 2\sqrt{\alpha \theta} \left\| \mathbf{S} x^{0} \right\|_{M} \left\| \mathbf{S} z^{0} \right\|_{M} - \frac{\alpha}{\theta} \left\| \mathbf{v} \right\|_{M}^{2} \right]$$

#### Abstract Proof.

First take a full expectation on *One-step inequality* with  $z^k, x^k$ . With the next two lemmas, we can prove that A is an uniform upper bound of full expectation on the last two terms of *One-step inequality*:

$$\theta\left(\theta\beta-\alpha\right)\left\|\mathbf{S}x^{k-1}\right\|_{M}^{2}+\alpha\theta\left(1-\alpha\theta\right)\left[2\left\langle\mathbf{S}x^{k-1},\mathbf{S}z^{k-1}\right\rangle_{M}-\left\|\mathbf{S}z^{k-1}\right\|_{M}^{2}\right].$$

Thus, we get

$$\mathbb{E}\left[\left\|x^{k}-z^{k}\right\|_{M}^{2}\right] \leq \mathbb{E}\left[\left\|x^{k-1}-z^{k-1}\right\|_{M}^{2}\right] + A.$$

Apply above inequality repeatedly, divide by  $k^2$  to conclude the proof.

**Remark.** To prove the main theorem of  $L^2$  convergence, apply Pazy's theorem on  $z^k$  to conclude the proof.

# **Proof of** $L^2$ convergence

# First, let's bound $||Sz^k||_M$ independent from k.

#### Lemma

 $\mathbb{T}: \mathcal{H} \to \mathcal{H}$  is a  $\theta$ -averaged with  $\theta \in (0, 1]$  and  $\mathbb{S} = \theta^{-1}(\mathbb{I} - \mathbb{T})$ . Let  $z^0, z^1, z^2, \cdots$  to be the iterates of (RC-FPI) with starting point  $z^0$ . Then,

$$\left\| \mathbf{S} z^{k} \right\|_{\mathcal{M}} \leq \left\| \mathbf{S} z^{k-1} \right\|_{\mathcal{M}} \leq \cdots \leq \left\| \mathbf{S} z^{0} \right\|_{\mathcal{M}}.$$

#### Proof.

With  $\mathbb{T}z - z = -\theta Sz$  and S being (1/2)-cocoercive operator,

$$\theta \left\langle \mathbb{ST}z - \mathbb{S}z, -\mathbb{S}z - \mathbb{ST}z \right\rangle_M \ge (1-\theta) \left\| \mathbb{ST}z - \mathbb{S}z \right\|_M^2 \ge 0,$$

Thus  $\|\mathbb{ST}z\|_M \leq \|\mathbb{S}z\|_M$  holds for any  $z \in \mathcal{H}$ , concluding the proof.

# Next, let's bound $\|\mathbb{E}[Sx^k]\|_M$ independent from k.

#### Lemma

 $\mathbb{T}: \mathcal{H} \to \mathcal{H}$  is a  $\theta$ -averaged with  $\theta \in (0, 1]$  and  $\mathbb{S} = \theta^{-1}(\mathbb{I} - \mathbb{T})$ . Let  $x^0, x^1, x^2, \cdots$  to be the iterates of (RC-FPI) with starting point  $x^0$ . Then,

$$\left\|\mathbb{E}\left[\mathbb{ST}_{\mathcal{I}^{k}}\ldots\mathbb{T}_{\mathcal{I}^{0}}x^{0}\right]\right\|_{M}\leq\beta^{1/2}\alpha^{-1}\left\|\mathbb{S}x^{0}\right\|_{M}$$

holds if  $\mathcal{I}^0, \mathcal{I}^1, \ldots, \mathcal{I}^k$  follow IID distribution with the uniform expected step-size condition with  $\alpha \in (0, 1]$  and  $\beta$  satisfies  $\beta \leq \alpha/\theta$ .

# **Proof of** *L*<sup>2</sup> **convergence**

#### Proof.

With the same distribution of  $\mathcal{I}$ , it is possible to show

$$\mathbb{E}_{\mathcal{I},X,Y}\left[\left\|\mathbb{T}_{\mathcal{I}}X - \mathbb{T}_{\mathcal{I}}Y\right\|_{M}^{2}\right] \leq \mathbb{E}_{X,Y}\left[\left\|X - Y\right\|_{M}^{2}\right].$$

for any random variable X, Y on  $\mathcal{H}$ . (::  $\beta \leq \alpha/\theta$  and (1/2)-cocoersivity.) Apply above inequality repeatedly and use Jensen's inequality:

$$\left\|\mathbb{E}\left[\mathbb{T}_{\mathcal{I}^{k}}\dots\mathbb{T}_{\mathcal{I}^{1}}X-\mathbb{T}_{\mathcal{I}^{k}}\dots\mathbb{T}_{\mathcal{I}^{1}}Y\right]\right\|_{\mathcal{M}}^{2}\leq\mathbb{E}\left[\left\|X-Y\right\|_{\mathcal{M}}^{2}\right]$$

Now set up X, Y as  $X = \mathbb{T}_{\mathcal{I}^0} x^0$ ,  $Y = x^0$ . Shift the index using IID,

$$\left\|\alpha \mathbb{E}\left[\theta \$ x^{k}\right]\right\|_{M} = \left\|\mathbb{E}\left[(\mathbb{T}_{\mathcal{I}^{k}} - \mathbb{I})\mathbb{T}_{\mathcal{I}^{k-1}} \dots \mathbb{T}_{\mathcal{I}^{0}} x^{0}\right]\right\|_{M} \leq \beta^{1/2} \left\|\theta \$ x^{0}\right\|_{M}.$$

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#### Theorem (Almost sure convergence of normalized iterate)

Let  $\mathbb{T}: \mathcal{H} \to \mathcal{H}$  be  $\theta$ -averaged with respect to  $\|\cdot\|$ -norm with  $\theta \in (0, 1)$ . Assume  $\mathcal{I}^0, \mathcal{I}^1, \ldots$  is sampled IID from a distribution satisfying the uniform expected step-size condition. Let  $x^0, x^1, x^2, \ldots$  be the iterates of (RC-FPI). Then

$$\frac{x^k}{k} \stackrel{\text{a.s.}}{\to} -\alpha \mathbf{v}$$

as  $k \to \infty$ .

**Remark.** We will prove for the *M*-norm with the assumption:

 $\alpha > \theta \beta$ .

As mentioned, we will alternatively prove the following result.

#### Lemma

Let  $\mathbb{T}: \mathcal{H} \to \mathcal{H}$  be  $\theta$ -averaged with respect to  $\|\cdot\|_M$  with  $\theta \in (0, 1]$ . Assume  $\mathcal{I}^0, \mathcal{I}^1, \ldots$  is sampled IID from a distribution satisfying the uniform expected step-size condition and  $\beta$  satisfies that  $\beta < \alpha/\theta$ . Let  $x^0, x^1, x^2, \ldots$  be the iterates of (RC-FPI). Then,  $x^k/k$  is strongly convergent to  $-\alpha \mathbf{v}$  in probability 1, i.e.

$$\frac{x^k}{k} \stackrel{\text{a.s.}}{\to} -\alpha \mathbf{v}$$

as  $k \to \infty$ , where **v** is the infimal displacement vector of  $\mathbb{T}$ .

#### Lemma (Robbins-Siegmund quasi-martingale theorem)

 $\mathcal{F}_1 \subset \mathcal{F}_2 \subset \ldots$  is a sequence of sub- $\sigma$ -algebras of  $\mathcal{F}$  where  $(\Omega, \mathcal{F}, P)$  is a probability space. When  $X_k, b_k, \tau_k, \zeta_k$  are non-negative  $\mathcal{F}_k$ -random variables such that

$$\mathbb{E}\left[X_{k+1} \mid \mathcal{F}_k\right] \leq \left(1 + b_k\right) X_k + \tau_k - \zeta_k,$$

 $\begin{array}{l} \lim_{k\to\infty} X_k \text{ exists and is finite and } \sum_{k=1}^{\infty} \zeta_k < \infty \text{ almost surely if} \\ \sum_{k=1}^{\infty} b_k < \infty, \sum_{k=1}^{\infty} \tau_k < \infty. \end{array}$ 

#### Abstract Proof.

The idea is to apply almost supermartingale theorem.

To do this, we first need to make an upper bound of the last two terms in *One-step inequality* without taking full expectation.

$$-\alpha\theta(1-\alpha\theta) \|\mathbf{S}x-\mathbf{S}z\|_{M}^{2}+\theta^{2}(\beta-\alpha^{2}) \|\mathbf{S}x\|_{M}^{2}$$

$$=\underbrace{-\theta(\alpha-\beta\theta) \|\mathbf{S}x-\frac{\alpha-\alpha^{2}\theta}{\alpha-\beta\theta}\mathbf{S}z\|_{M}^{2}}_{\leq 0} +\underbrace{\frac{\alpha\theta^{2}(1-\alpha\theta)(\beta-\alpha^{2})}{\alpha-\beta\theta}}_{=:B\geq 0} \|\mathbf{S}z\|_{M}^{2}$$

Now consider a sequence  $z^0, z^1, z^2, \cdots$  of (FPI by  $\overline{\mathbb{T}}$ ).

#### Abstract Proof continued.

From  $\left\| \mathbb{S}z^{k} \right\|_{M} \leq \left\| \mathbb{S}z^{0} \right\|_{M}$ ,

$$\mathbb{E}_{\mathcal{I}^{k}}\left[\left\|\frac{x^{k}}{k}-\frac{z^{k}}{k}\right\|_{M}^{2}\mid\mathcal{F}_{k-1}\right] \leq \left\|\frac{x^{k-1}}{k-1}-\frac{z^{k-1}}{k-1}\right\|_{M}^{2}+\frac{B}{k^{2}}\left\|\$z^{0}\right\|_{M}^{2}.$$

Since  $\sum \frac{B}{k^2} \|\mathbf{S}z^0\|_M^2 < \infty$ , apply the R-S quasi-martingale theorem. Then,  $\|\frac{x^k}{k} - \frac{z^k}{k}\|_M^2$  converges almost surely to some random variable. With Fatou's lemma and  $L^2$  convergence,

$$\mathbb{E}\left[\lim_{k\to\infty} \left\|\frac{x^{k}}{k} - \frac{z^{k}}{k}\right\|_{M}^{2}\right] \leq \lim_{k\to\infty} \mathbb{E}\left[\left\|\frac{x^{k}}{k} - \frac{z^{k}}{k}\right\|_{M}^{2}\right] = 0.$$
  
Thus, with probability 1,  $\lim_{k\to\infty} \left\|\frac{x^{k}}{k} - \frac{z^{k}}{k}\right\|_{M}^{2} = 0.$ 

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# Variance of Normalized iterates

Goal. Here's the main result we will prove in this section.

#### Theorem

Let  $\mathbb{T}: \mathcal{H} \to \mathcal{H}$  be  $\theta$ -averaged with respect to  $\|\cdot\|_M$  with  $\theta \in (0, 1]$ . Assume  $\mathcal{I}^0, \mathcal{I}^1, \ldots$  is sampled IID from a distribution satisfying the uniform expected step-size condition and  $\beta$  satisfies that  $\beta < \alpha/\theta$ . Let  $x^0, x^1, x^2, \ldots$  be the iterates of (RC-FPI).

(a) If 
$$\mathbf{v} \in \operatorname{range}(\mathbb{I} - \mathbb{T})$$
, then

$$\limsup_{k \to \infty} k \mathbb{E} \left[ \left\| \frac{x^k}{k} + \alpha \mathbf{v} \right\|_M^2 \right] \le (\beta - \alpha^2) \left\| \mathbf{v} \right\|_M^2.$$

(b) In general, regardless of whether  $\mathbf{v}$  is in range  $(\mathbb{I} - \mathbb{T})$  or not,

$$\limsup_{k \to \infty} k \operatorname{Var}_{M} \left( \frac{x^{k}}{k} \right) \leq (\beta - \alpha^{2}) \left\| \mathbf{v} \right\|_{M}^{2}.$$

**Example.** Before moving on to the proofs, here's check some examples.

**Ex 1. Equality holds:** Consider the translation operator  $\mathbb{T}(x) = x - \mathbf{v}$ . Choose the norm as  $\|\cdot\|$  and consider the distribution of  $\mathcal{I}$  as an uniform distribution on standard basis.

When  $x^0, x^1, x^2, \ldots$  are the iterates of (RC-FPI) with  $\mathbb{T}$ , then

$$k \operatorname{Var}\left(\frac{x^{k}}{k}\right) = \alpha \left(1 - \alpha\right) \|\mathbf{v}\|^{2}$$

for k = 1, 2, ..., and the variance bound holds with equality.

**Remark.** In this scenario, each step is independent to each other. The result is identical to the result of *Central Limit Theorem*.

### Variance of Normalized iterates

Ex 2. Strict inequality Define  $\mathbb{T} \colon \mathbb{R}^2 \to \mathbb{R}^2$  as  $\mathbb{T} = \mathbb{I} - \operatorname{Proj}_{x+y=1}$ , or

$$\mathbb{T}: (x,y) \mapsto \left(x - \frac{1+x-y}{2}, y - \frac{1+y-x}{2}\right),$$

which is 1/2-averaged and has the infimal displacement vector (1/2, 1/2). When  $(x^0, y^0), (x^1, y^1), (x^2, y^2), \ldots$  are the iterates of (RC-FPI) with T,

$$\limsup_{k\to\infty} k \operatorname{Var}_M\left(\frac{(x^k, y^k)}{k}\right) = \frac{1}{24}.$$

On the other hand, the right hand side of the inequality is

$$\alpha (1 - \alpha) \|\mathbf{v}\|^2 = \frac{1}{2} \left(1 - \frac{1}{2}\right) \|\mathbf{v}\|^2 = \frac{1}{8}$$

**Remark.** This is also an example that  $\lim_{k\to\infty} \mathbb{E}\left[\left\|\theta \$ x^k\right\|_M^2\right] \neq \|\mathbf{v}\|_M^2$ .

# Variance of Normalized iterates



Simulation of prior example, 15 iterations, ran 10 times.

# Idea of proof

### Idea of proof.

First generate a sequence  $z^0, z^1, z^2, \cdots$  as (FPI by  $\overline{\mathbb{T}}$ ) with  $z^0 = z$ . From the *One-step inequality*, with full expectation,

$$\mathbb{E}\left[k\left\|\frac{x^{k}}{k}-\frac{z^{k}}{k}\right\|_{M}^{2}\right] \leq \frac{1}{k}\left\|x^{0}-z^{0}\right\|_{M}^{2} + \mathbb{E}\left[\frac{1}{k}\sum_{j=0}^{k-1}U^{j}\right]$$

where  $U^0, U^1, U^2, \ldots$  is a sequence of random variables:

$$U^{k} = -\alpha \left(\theta^{-1} - \alpha\right) \left\|\theta \mathbf{S} x^{k} - \theta \mathbf{S} z^{k}\right\|_{M}^{2} + \theta^{2} \left(\beta - \alpha^{2}\right) \left\|\mathbf{S} x^{k}\right\|_{M}^{2}$$

The key of this proof is to bound  $U^k$  asymptotically as  $k \to \infty$ ,

$$U^k \lesssim (\beta - \alpha^2) \left\| \mathbf{v} \right\|_M^2.$$

Then due to Cesàro mean we may conclude the proof.

#### Abstract Proof.

- 1. Choose z with  $\|\mathbf{v}\|_M \le \|\theta \mathbf{S} z^k\|_M \le \|\theta \mathbf{S} z\|_M \le \|\mathbf{v}\|_M + \epsilon$ . From convexity of range set, this gives  $\mathbf{v} \simeq \theta \mathbf{S} z^k \simeq \theta \mathbf{S} z$ .
- 2.  $x^k z$  and z are nearly orthogonal for after sufficient iteration:

$$\left\|\theta \mathbf{S} x^{k}\right\|_{M}^{2} \simeq \left\|\theta \mathbf{S} x^{k} - \theta \mathbf{S} z\right\|_{M}^{2} + \left\|\theta \mathbf{S} z\right\|_{M}^{2}.$$

Nearly orthogonal property will be handled in the next lemma.

3. Then we use this to build an inequality of a form:

$$U^{k} \lesssim \underbrace{-\theta^{-1} \left(\alpha - \beta \theta\right) \left\|\theta \$ x^{k} - \mathbf{v}\right\|_{M}^{2}}_{\leq 0} + \underbrace{\left(\beta - \alpha^{2}\right) \left\|\mathbf{v}\right\|_{M}^{2}}_{\text{RHS of the theorem}},$$
  
as  $k \to \infty$ .

# **Proof of Variance of Normalized iterates**

#### Abstract Proof.

To be precise, with probability 1, when  $x^0, x^1, x^2, \cdots$  are generated, for  $\delta \in (0, \pi/2)$  and  $z \in \mathcal{H}$ , there exists  $N_{\delta,z}$  s.t. for all  $k > N_{\delta,z}$ ,

$$U^{k} \leq \theta^{2} \left(\beta - \alpha^{2}\right) \left\| \mathbf{S} z \right\|_{M}^{2} + \frac{\theta}{\alpha - \beta \theta} \tilde{\tau}_{\delta, z}^{2},$$

where  $\tilde{\tau}_{\delta,z}$  depends only on  $\delta, z$  and  $\tilde{\tau}_{\delta,z} \to 0$  as  $\delta \to 0$  and  $\theta \$ z \to \mathbf{v}$ .

1. Since  $N_{\delta,z}$  also depends on  $\{x^k\}$ , apply Cesàro mean:

$$\limsup_{k \to \infty} \left[ \frac{1}{k} \sum_{j=0}^{k-1} U^j \right] \le \theta^2 \left( \beta - \alpha^2 \right) \left\| \mathbf{\$} z \right\|_M^2 + \frac{\theta}{\alpha - \beta \theta} \tilde{\tau}_{\delta, z}^2$$

- 2. Apply expectation, then use Reverse Fatou's lemma (since  $U^j \leq B \|Sz\|_M^2$  holds) to bound the variance term.
- 3. Finally, take the limit  $\delta \to 0$  and  $\theta Sz \to \mathbf{v}$  to conclude the proof.

### Lemma (Nearly orthogonal)

Let  $\mathbb{T}: \mathcal{H} \to \mathcal{H}$  be  $\theta$ -averaged with respect to  $\|\cdot\|_M$  with  $\theta \in (0, 1]$ . Consider a sequence  $y^0, y^1, y^2, \ldots$  in  $\mathcal{H}$  such that its normalized iterate converges strongly to  $-\gamma \mathbf{v}$  for some  $\gamma > 0$ :

$$\lim_{k\to\infty}\frac{y^k}{k}=-\gamma\mathbf{v}.$$

Then, for any  $\delta \in (0, \pi/2)$  and  $z \in \mathcal{H}$ ,  $\exists N_{\delta, z} \in \mathbb{N}$  s.t.,  $\forall k > N_{\delta, z}$ ,

$$\left\langle \mathbf{v}, \mathbf{S}y^{k} - \mathbf{S}z \right\rangle_{M} \leq \left\| \mathbf{v} \right\|_{M} \left\| \mathbf{S}y^{k} - \mathbf{S}z \right\|_{M} \sin \delta.$$

The proof is quite technical. We replace the proof with a figure describing the dynamics, while the full proof is available in the paper.

# Nearly orthogonal

Idea of proof.



# Relation between variance and range set

**Remark.** Throughout observations, we could check that there was a weak relation between the shape of the range set range  $(\mathbb{I} - \mathbb{T})$  near **v** and the variance value.



We interpret such phenomenon happen when the following inequality can be non-tight even after sufficient iterations.

$$-\theta^{-1} \left(\alpha - \beta \theta\right) \left\|\theta \$ x^k - \mathbf{v}\right\|_M^2 \le 0$$

Introduction

RC-FPI

Convergence of Normalized iterates L<sup>2</sup> convergence of normalized iterate Almost sure convergence of normalized iterate

Variance of Normalized iterates

Infeasibility Detection

Extension to Nonorthogonal basis

### Goal. Infeasibility detection method for (RC-FPI) by hypothesis testing.

#### Theorem

Let  $\mathbb{T}: \mathcal{H} \to \mathcal{H}$  be  $\theta$ -averaged with respect to  $\|\cdot\|_M$  with  $\theta \in (0, 1]$ . Consider a null hypothesis of  $\|\mathbf{v}\|_M \leq \delta$  with small  $\delta$  satisfying  $\alpha \delta < \epsilon$ . Assume  $\mathcal{I}^0, \mathcal{I}^1, \ldots$  is sampled IID from a distribution satisfying the uniform expected step-size condition and  $\beta$  satisfies that  $\beta < \alpha/\theta$ . Let  $x^0, x^1, x^2, \ldots$  be the iterates of (RC-FPI). Then

$$\mathbb{P}\left(\left\|\frac{x^{k}}{k}\right\|_{M} \ge \varepsilon\right) \lesssim \frac{\left(\beta - \alpha^{2}\right)\delta^{2}}{k(\varepsilon - \alpha\delta)^{2}}$$

as  $k \to \infty$ , where **v** is the infimal displacement vector of  $\mathbb{T}$ .

Therefore, for any statistical significance level  $p \in (0, 1)$ , the test

$$\left\|\frac{x^k}{k}\right\|_M \ge \varepsilon$$

with

$$k \gtrsim rac{\left(eta - lpha^2
ight)\delta^2}{p\left(arepsilon - lpha\delta
ight)^2}$$

can reject the null hypothesis and conclude that  $\|\mathbf{v}\|_M > \delta$ , which implies that the problem is inconsistent.

#### Proof.

Here, we provide a simpler case where  $\mathbf{v} \in \operatorname{range}(\mathbb{I} - \mathbb{T})$ .

By the triangle inequality, Markov inequality, and the variance theorem we just proved, under the null hypothesis,

$$\mathbb{P}\left(\left\|\frac{\mathbf{x}^{k}}{k}\right\|_{M} \ge \varepsilon\right) \le \mathbb{P}\left(\left\|\frac{\mathbf{x}^{k}}{k} + \alpha \mathbf{v}\right\|_{M} \ge \varepsilon - \alpha\delta\right)$$
$$\le \frac{1}{(\varepsilon - \alpha\delta)^{2}} \mathbb{E}\left[\left\|\frac{\mathbf{x}^{k}}{k} + \alpha \mathbf{v}\right\|_{M}^{2}\right]$$
$$\le \frac{(\beta - \alpha^{2})\delta^{2}}{k(\varepsilon - \alpha\delta)^{2}},$$

as  $k \to \infty$ . The full proof use the inequality about the variance.

Introduction

RC-FPI

Convergence of Normalized iterates *L*<sup>2</sup> convergence of normalized iterate Almost sure convergence of normalized iterate

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Extension to Nonorthogonal basis

### Review

### Review.

In general  $\|\cdot\|_M$ -norm,

• if  $\alpha \geq \theta \beta$ ,

$$\frac{\mathbf{x}^{k}}{k} \stackrel{L^{2}}{\to} -\alpha \mathbf{v}.$$

• If 
$$\alpha > \theta \beta$$
,  

$$\frac{x^k}{k} \stackrel{\text{a.s.}}{\to} -\alpha \mathbf{v},$$
and  $\limsup_{k \to \infty} k \operatorname{Var}_M \left(\frac{x^k}{k}\right) \le (\beta - \alpha^2) \|\mathbf{v}\|_M^2.$ 

When we use  $\|\cdot\|$ -norm, we can choose  $\beta = \alpha$ . Thus the conditions  $\alpha \ge \theta \beta$  and  $\alpha > \theta \beta$  are replaced with  $\theta \in (0, 1]$  and  $\theta \in (0, 1)$ .

**Remark.** However, optimization methods such as (PG-EXTRA) uses averaged operator that's 1-Lipschitz in  $\|\cdot\|_M$ -norm with  $M \neq \mathbb{I}$ .

Now let's consider an update with nonorthogonal basis.

• The underlying space  ${\mathcal H}$  with extra  ${\mathcal H}_0$  block, making  ${\mathcal H}$  as

$$\mathcal{H} = \mathcal{H}_0 \oplus \mathcal{H}_1 \oplus \mathcal{H}_2 \oplus \ldots \mathcal{H}_m$$

• Consider two subspaces  $U_1$  and  $U_2$  of  $\mathcal{H}$  as

 $U_1 = \mathcal{H}_0 \times 0 \times 0 \times \cdots \times 0, \quad U_2 = 0 \times \mathcal{H}_1 \times \mathcal{H}_2 \times \cdots \times \mathcal{H}_m.$ 

- We assume that with *M*-inner product of  $\mathcal{H}$ ,
  - Block components in  $U_2$  are orthogonal.
  - $U_1$  and  $U_2$  need not be orthogonal.

**Question.** When are the conditions  $\alpha \ge \theta\beta$  and  $\alpha > \theta\beta$  satisfied?

### Definition (Friedrichs angle)

The cosine of the *Friedrichs angle*  $c_F$  between two subspaces  $U_1$  and  $U_2$  is defined as a smallest value among  $c \leq 1$  such that satisfies:

 $|\langle u_1, u_2 \rangle_M| \le c \|u_1\|_M \|u_2\|_M \quad \forall u_1 \in U_1, u_2 \in U_2.$ 

**Remark.** When  $M = \mathbb{I}$ ,  $c_F = 0$ .

#### Lemma

Suppose the subspaces  $U_1, U_2$  of  $\mathcal{H}$  with  $U_1 \cap U_2 = \{0\}$  satisfy:

$$|\langle u_1, u_2 \rangle_M| \le c_F ||u_1||_M ||u_2||_M, \quad c_F \le \sqrt{\frac{1-\theta}{1-\alpha\theta}}$$

for any  $u_1 \in U_1, u_2 \in U_2$ . Then, there exists  $\beta \ge 0$  such that  $\beta \theta \le \alpha$  and

$$\mathbb{E}_{\mathcal{I}}\left[u_{\mathcal{I}}\right] = \alpha u, \quad \mathbb{E}_{\mathcal{I}}\left[\left\|u_{\mathcal{I}}\right\|_{\mathcal{M}}^{2}\right] \leq \beta \left\|u\right\|_{\mathcal{M}}^{2}.$$

If  $c_{\mathsf{F}} < \sqrt{\frac{1-\theta}{1-\alpha\theta}}$ , then there exists  $\beta$  with  $\beta\theta < \alpha$ .

**Remark.** Setting  $\beta$  as  $\beta = \alpha^2 + \frac{\alpha - \alpha^2}{1 - c_F^2}$  proves the lemma.

Thus, in this case of (RC-FPI),

#### Extension to nonorthogonal basis.

In general  $\|\cdot\|_{M}$ -norm, • if  $c_{F} \leq \sqrt{\frac{1-\theta}{1-\alpha\theta}}$ ,  $\frac{x^{k}}{k} \xrightarrow{L^{2}} -\alpha \mathbf{v}$ . • If  $c_{F} < \sqrt{\frac{1-\theta}{1-\alpha\theta}}$ ,  $\frac{x^{k}}{k} \xrightarrow{\text{a.s.}} -\alpha \mathbf{v}$ , and  $\limsup_{k\to\infty} k \operatorname{Var}_{M}\left(\frac{x^{k}}{k}\right) \leq (\beta - \alpha^{2}) \|\mathbf{v}\|_{M}^{2}$ .

**Remark.**  $\beta$  value is guaranteed to be at most  $\beta = \alpha^2 + \frac{\alpha - \alpha^2}{1 - c_F^2}$ .

### Definition (PG-EXTRA)

Consider the convex optimization problem

$$\underset{x \in \mathbb{R}^d}{\text{minimize}} \quad \sum_{i=1}^m f_i(x),$$

where  $f_i : \mathbb{R}^d \to \mathbb{R}$  is closed, convex, and proper function. The decentralized algorithm *PG-EXTRA* is a iterative solver:

$$x_{i}^{k+1} = \operatorname{Prox}_{\tau f_{i}} \left( \sum_{j=1}^{m} W_{ij} x_{i}^{k} - w_{i}^{k} \right)$$
  

$$w_{i}^{k+1} = w_{i}^{k} + \frac{1}{2} \left( x_{i}^{k} - \sum_{j=1}^{m} W_{ij} x_{j}^{k} \right)$$
(PG-EXTRA)

**Remark.** When using (PG-EXTRA), it is computed with *m* agents. *W* is a matrix representing the connection between agents, satisfying:

- W is symmetric m by m matrix.
- $W_{ij} = 0$  if  $i \neq j$  and agents *i* and *j* are not connected.
- $N(I W) = \operatorname{span}(1)$  and  $1 = \lambda_1 > \max\{|\lambda_2|, \cdots, |\lambda_m|\}.$

Each agent *i* handles  $x_i$  and  $w_i$  value. At each iteration, only communication between connected agents are required.

# Application in PG-EXTRA

A randomized coordinate-update version of (PG-EXTRA) performs:

- randomly chooses *i* among  $1, 2, \ldots, m$  to update  $x_i^k$ ,
- every  $w_1, w_2, \ldots, w_m$  gets updated at each iterations.

### RC-PG-EXTRA.

while Not converged do Sample:  $\mathcal{I}$ **for** *i* such that  $\mathcal{I}_i \neq 0$  **do**  $\Delta x_i = \operatorname{Prox}_{\tau f} \left( [W_X]_i - w_i \right) - x_i$ **Update:**  $x_i = x_i + \mathcal{I}_i \Delta x_i$ for  $j \in N_i \cup \{i\}$  do **Send:**  $\Delta x_i$  From *i*th agent to *j*th agent.  $[W_X]_i = [W_X]_i + W_{ii}\Delta x_i$ end for end for end while

#### Corollary

Assume  $\mathcal{I}^0, \mathcal{I}^1, \ldots$  is sampled IID from a distribution satisfying the uniform expected step-size condition. Perform (RC-PG-EXTRA). If the minimum eigenvalue of the mixing matrix  $W \in \mathbb{R}^m$  satisfies:

$$\lambda_{\min}(W) > -\frac{\alpha}{2-\alpha},$$

then the results of previous theorems holds:

$$\frac{x^{k}}{k} \stackrel{\text{a.s.}}{\to} -\alpha \mathbf{v}, \quad \limsup_{k \to \infty} k \operatorname{Var}_{M} \left( \frac{x^{k}}{k} \right) \leq (\beta - \alpha^{2}) \left\| \mathbf{v} \right\|_{M}^{2}.$$

Additionally, here is the infimal displacement vector of (PG-EXTRA).

#### Lemma

The infimal displacement vector  $\mathbf{v} = (\mathbf{v}_1, \dots, \mathbf{v}_m)$  of (PG-EXTRA) is

$$\mathbf{v}_{i} = \begin{bmatrix} \frac{\tau}{m} \sum_{j=1}^{m} g_{j} \\ -\frac{1}{2} \left( y_{i} - \sum_{j=1}^{m} W_{ij} y_{j} \right) \end{bmatrix}$$

for  $i = 1, \ldots, m$ , where  $(y_1, y_2, \ldots, y_m)$  and  $(g_1, g_2, \ldots, g_m)$  are

$$\operatorname{argmin}_{\substack{y_1, y_2, \dots, y_m \in \mathbb{R}^d \\ g_j \in \partial f_j(y_j), 1 \le j \le m}} \left\| \frac{\tau}{m} \sum_{j=1}^m g_j \right\|^2 + \frac{1}{2} \sum_{i,j=1}^m W_{ij} \left\| y_i - y_j \right\|^2.$$

# **Application in PG-EXTRA**

**Remark.** We observed faster convergence to the infimal displacement vector when randomized, in the inconsistent case of (PG-EXTRA). Here, the notion of *faster* regards on the communication count, not iterations.



(Left) Network used in our experiment, consisting of m = 40 agents, with agents 2,..., 40 densely connected. (Right) Graph of  $\|(\mathbf{x}^k, \mathbf{u}^k)/k + \alpha \mathbf{v}\|^2$  against the communication count for (PG-EXTRA) and (RC-PG-EXTRA).

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